

Yang & Barron.

Minimax Estimation:

$$\theta \in S \subseteq \Theta$$

"parameter", could be functions, densities, etc.

Loss function: $d(\theta, \hat{\theta})$

$\hat{\theta}$ = estimator of θ , $(x_1, \dots, x_n) \mapsto \hat{\theta}$.

$$r_n = \min_{\hat{\theta}} \max_{\theta \in S} \mathbb{E}[d^2(\theta, \hat{\theta})] : \text{function of } S \text{ and } n, r_n = r_n(S).$$

Examples: (non-parametric regression)

$$y_i = f(x_i) + \varepsilon_i, \quad f: [0, 1] \rightarrow \mathbb{R}, \quad i=1, \dots, n, \quad \varepsilon_i \sim \mathcal{N}(0, \sigma^2)$$

Data: $(x_i, y_i), i=1, \dots, n$

$\theta \equiv f$, $S^\alpha = \{f \in C^\alpha \mid f \text{ is } \alpha\text{-times cts differentiable, } \|f\| \leq K\}$ (Sobolev-ball)

$$r_n(S^\alpha) \asymp \left(\frac{1}{n}\right)^{\frac{2\alpha}{2\alpha+1}} \quad d^2(f, \hat{f}) = \int_0^1 |f(t) - \hat{f}(t)|^2 dt.$$

$$S^k \subseteq S^{k-1} \subseteq \dots \subseteq S^1.$$

Need to measure complexity of parameter space.

ε -packing

Finite set $N_\varepsilon \subset S$ s.t. $\forall \theta, \theta' \in N_\varepsilon, \theta \neq \theta' \quad d(\theta, \theta') > \varepsilon$.

packing ε -entropy $\log |N_\varepsilon|$

Examples:

(a) $\{\theta \in \mathbb{R}^d \mid \|\theta\|_2 \leq 1\}$.

pack it with Euclidean balls. (last semester).

$$M(\varepsilon) = \log |N_\varepsilon| = d \log\left(\frac{1}{\varepsilon}\right)$$

(b) Sobolev-ball S^α (Birman, 1967)

d -dimension $\alpha = \text{smoothness}$.

$$\log |N_\varepsilon| \asymp \left(\frac{1}{\varepsilon}\right)^{d/2}. \quad (\text{also for covering})$$

ϵ -covering:

A set $G_\epsilon \subset S$ is an ϵ -cover of S if
 $\forall \theta \in S, \exists \text{ some } \theta' \in G_\epsilon \text{ st. } d(\theta, \theta') < \epsilon$.

$$\underbrace{\log |G_\epsilon|}_{V(\epsilon)} = \text{covering entropy}$$

Step 1: choose $\epsilon_n \rightarrow 0$ st. $\epsilon_n^2 n = V(\epsilon_n)$

Step 2: choose $\epsilon_{nd} \rightarrow 0$ st
 $M(\epsilon_{nd}) = 4n\epsilon_n^2 + 2\log 2$

Thm 1: $\min_{\hat{\theta}} \max_{\theta \in S} \mathbb{E}[d^2(\theta, \hat{\theta})] \geq (\text{const}) \epsilon_{nd}^2$

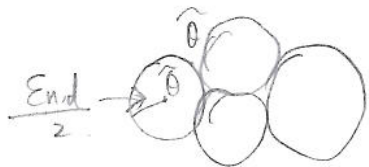
Example: $\epsilon_n^2 n = \left(\frac{1}{\epsilon_n}\right)^{d/2} \Rightarrow \epsilon_n^2 = \left(\frac{1}{n}\right)^{\frac{2d}{2d+d}}$
 $d=1, 2 \Rightarrow n^{-\frac{4}{5}}$

proof: Construct $N_{\epsilon_{nd}}$ packing set (with $|N_{\epsilon_{nd}}|$ elements)

Let $\hat{\theta}$ be any estimator.

$$\text{set } \tilde{\theta} = \underset{\theta' \in N_{\epsilon_{nd}}}{\text{argmin}} d(\theta', \hat{\theta}).$$

claim: $\forall \theta \in N_{\epsilon_{nd}}, \theta \neq \tilde{\theta}, d(\theta, \hat{\theta}) \geq A \epsilon_{nd}/2$.



$$\min_{\hat{\theta}} \max_{\theta \in S} \mathbb{P}_\theta (d(\theta, \hat{\theta}) \geq \frac{A}{2} \epsilon_{nd}).$$

$$\geq \min_{\hat{\theta}} \max_{\theta \in N_{\epsilon_{nd}}} \mathbb{P}_\theta (d(\theta, \hat{\theta}) \geq \frac{A}{2} \epsilon_{nd}).$$

$$= \min_{\hat{\theta}} \max_{\theta \in N_{\epsilon_{nd}}} \mathbb{P}_\theta (\theta \neq \tilde{\theta})$$

$\tilde{\theta}$ is implicitly defined through $\hat{\theta}$.

$$\geq \min_{\hat{\theta}} \sum_{\theta \in N_{\epsilon_{nd}}} w(\theta) \mathbb{P}_\theta (\theta \neq \tilde{\theta}) = \min_{\hat{\theta}} \mathbb{P}_w (\theta \neq \tilde{\theta})$$

$w(\theta) \equiv$ uniform over N_{end} (parameter space)

"codewords" \equiv ball centers $\theta \in N_{\text{end}}$.

"channel" given $\theta \rightarrow (x_1, \dots, x_n)$.

Fano's \leq : Estimate $\theta \in N_{\text{end}}$ from X

$$P_{\text{err}} \geq \frac{H(\theta|X) - 1}{\log |N_{\text{end}}|}.$$

our case: θ is uniform over N_{end} , $H(\theta) = \log |N_{\text{end}}|$.

$$I(\theta; X) = H(\theta) - H(\theta|X).$$

$$P_w(\theta \neq \hat{\theta}) \geq 1 - \frac{I(\theta; X_1^n)}{\log |N_{\text{end}}|} + o(1)$$

Remains to upper bound

$$\begin{aligned} \text{(rew)} \quad I(\theta; X_1^n) &= D(P(\theta; X_1^n) \| P(\theta) P(X_1^n)) \\ &= \sum_{\theta} w(\theta) \int P(X_1^n | \theta) \log \frac{P(X_1^n | \theta)}{P_w(X_1^n)} dx \\ &\quad \underbrace{D(P(X_1^n | \theta) \| P_w(X_1^n))} \end{aligned}$$

$$\text{claim: } \leq \sum_{\theta} w(\theta) D(P(X^n | \theta) \| q(X^n))$$

for any choice of $q(\cdot)$ over $X \times X \times \dots \times X$ n -times.

$$\leq \max_{\theta \in N_{\text{end}}} D(P(X^n | \theta) \| q(X^n)).$$

Last step: construct a ϵ -covering in KL-divergence.

$G_{\epsilon n}$ For all $\theta \in S$, $\exists \bar{\theta} \in G_{\epsilon n}$, with

$$d_K(\theta, \bar{\theta}) \leq \epsilon_n^2$$

$$\text{choose } q(X^n) = \frac{1}{|G_{\epsilon n}|} \sum_{\theta \in G_{\epsilon n}} P(X^n | \theta).$$

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$$P_{\text{err}} \geq 1 - \frac{V(\bar{\epsilon}_n) + n\bar{\epsilon}_n^2}{M(\bar{\epsilon}_n, d) = 4n\bar{\epsilon}_n^2}$$

estimation $\xrightarrow{\text{packing}}$ hypothesis testing
unified view of minimax